

# Supplementary exercises for sheet 6

You don't have to hand out those supplementary exercises; if you struggle with one of them, Alix Deleporte (who wrote them) will be happy to help!

## Independence of real random variables

**Exercise 1.** Give a necessary and sufficient condition for a random variable  $X$  to be independent from itself.

**Exercise 2.** Let  $U$  and  $X$  be two independent random variables:  $U$  follows a uniform law on  $[0, 1]$  and  $X$  follows  $\mathcal{E}(1)$ .

1. Compute  $\mathbb{P}(\max(U, X) \leq t)$  for every  $t > 0$ .
2. Draw the density of  $\max(U, X)$ .

**Exercise 3.** Let  $(X_n)_{n \geq 1}$  be a sequence of real random variables on a probabilistic set  $(\Omega, \mathbb{F}, \mathbb{P})$ . Suppose that there exists a random variable  $X$  on  $(\Omega, \mathbb{F}, \mathbb{P})$  such that  $X_n \rightarrow X$  almost surely. (reminder: this means that for almost all  $\omega$ ,  $X_n(\omega) \rightarrow X(\omega)$ ).

For every  $\epsilon > 0$  and  $n \in \mathbb{N}$ , let

$$A_n^\epsilon = \{\omega \in \Omega, |X_n(\omega) - X(\omega)| < \epsilon\}.$$

Prove carefully that, for every  $\epsilon > 0$ , as  $n \rightarrow +\infty$ ,  $\mathbb{P}(A_n^\epsilon) \rightarrow 1$ .

**Exercise 4.** Let  $X_1, \dots, X_n$  be independent, identically distributed (iid) real random variables, with density  $f$ . We are interested in the random variable

$$Z = \max(X_1, \dots, X_n) - \min(X_1, \dots, X_n).$$

We let  $F : \mathbb{R} \rightarrow [0, 1]$  be the following function:  $F : t \mapsto \mathbb{P}(X_1 < t)$ .

1. Show that, for all  $u < v$ ,

$$\mathbb{P}(\min(X_1, \dots, X_n) \leq u, \max(X_1, \dots, X_n) < v) = (F(v) - F(u))^n.$$

Show that, for all  $t > 0$ ,

$$\mathbb{P}\left(\bigcap_{j=2}^n \{X_1 < X_j < X_1 + t\}\right) = \int_{\mathbb{R}} (F(x+t) - F(x))^{n-1} f(x) dx.$$

Show that, for all  $t > 0$ ,

$$\{Z < t\} = \bigcup_{i=1}^n \bigcap_{j \neq i} \{X_i < X_j < X_i + t\}$$

and that the union is disjoint.

2. Deduce an expression of  $t \mapsto \mathbb{P}(Z < t)$  involving  $f$  and  $F$ .
3. We now suppose that  $X_1, \dots, X_n$  follow a uniform distribution on  $[0, 1]$ . Make the above question explicit, show that  $Z$  admits a density, and determine this density.

**Exercise 5.** The Gamma function  $\Gamma : (0, +\infty) \rightarrow (0, +\infty)$  is defined as follows:  $\Gamma : a \mapsto \int_0^{+\infty} x^{a-1} e^{-x} dx$ .

Given  $a > 0$  and  $\lambda > 0$ , the gamma distribution  $\gamma_{a,\lambda}$  is defined by its density  $f_{a,\lambda} : \mathbb{R} \rightarrow \mathbb{R}^+$  relatively to the Lebesgue measure:

$$f_{a,\lambda} : x \mapsto \frac{\lambda^a}{\Gamma(a)} \exp(-\lambda x) x^{a-1} \mathbb{1}(x \geq 0).$$

1. Show that this is a probability distribution; determine its expectation.
2. Let  $V_1, \dots, V_n$  be iid random variables, with law  $\mathcal{E}(\lambda)$ . Show, by induction, that the law of  $V_1 + V_2 + \dots + V_n$  is  $\gamma_{n,\lambda}$ .
3. Let  $X$  and  $Y$  be two iid random variables with respective laws  $\gamma_{a,\lambda}$  and  $\gamma_{b,\lambda}$ . Show that  $X + Y$  follows the law  $\gamma_{a+b,\lambda}$ .
4. Let  $Z$  be a real random variable following  $\mathcal{N}(0, 1)$ . Show that  $Z^2$  follows  $\gamma(\frac{1}{2}, \frac{1}{2})$ . What is the distribution of the sum of  $n$  squares of iid standard Gaussian random variables?

**Exercise 6.** Let  $(X_n)_{n \geq 1}$  a sequence of independent random variables, taking values in  $[0, 1]$ . The goal of this exercise is to show that the following conditions are equivalent:

- (i)  $\sum_{n=1}^{\infty} \mathbb{E}[X_n] < \infty$ .
- (ii) almost surely,  $\sum_{n=1}^{+\infty} X_n < \infty$ .
- (iii) with non-zero probability,  $\sum_{n=1}^{+\infty} X_n < \infty$ .

For every  $n$ , we let  $S_n = X_1 + \dots + X_n$ , and we let  $S = \lim_{n \rightarrow +\infty} S_n \in [0, +\infty]$ . In particular, condition (ii) can be written  $\mathbb{P}(S < \infty) = 1$ ; clearly (ii) implies (iii).

1. Suppose (i) holds. Prove that  $\mathbb{E}\left[\sum_{n=1}^{+\infty} X_n\right] < \infty$ , then prove (ii).

In the next questions we assume that (iii) holds.

2. Show that  $\mathbb{E}[e^{-S}] > 0$ .
3. Show that, for all  $n \geq 1$ ,

$$\mathbb{E}[e^{-S_n}] = \prod_{k=1}^n \mathbb{E}[e^{-X_k}].$$

Deduce (with a correct argument) an expression for  $\mathbb{E}[e^{-S}]$ .

4. Show that

$$-\sum_{n=1}^{\infty} \ln(\mathbb{E}[e^{-X_n}]) < \infty.$$

5. Show that, for all  $x \in [0, 1]$ ,

$$e^{-x} \leq 1 - e^{-1}x \leq e^{-e^{-1}x}.$$

(Hint: use convexity).

6. Deduce that

$$\mathbb{E}[e^{-X_n}] \leq e^{-e^{-1}\mathbb{E}[X_n]}.$$

7. Conclude: (iii) implies (i)

## Strong convergence and Borel-Cantelli

**Exercise 7.** Give an example of sequence of events  $(A_n)_{n \geq 1}$  in a probability space  $\Omega$  such that  $\sum_{n=1}^{+\infty} \mathbb{P}(A_n) = +\infty$  but  $\mathbb{P}(\limsup(A_n)) = 0$ . (In particular, the  $A_n$ 's are not independent !)

**Exercise 8.** Let  $f : \mathbb{R} \rightarrow \mathbb{R}$  be bounded and such that  $\forall x, y, |f(x) - f(y)| \leq C|x - y|$ . Let  $X_n \rightarrow X$  a.s.

1. Show that  $\mathbb{E}[f(X_n)] - \mathbb{E}[f(X)] \rightarrow 0$ .
2. Let  $\epsilon > 0$ . Show that  $\mathbb{P}(|f(X_n) - f(X)| \geq \epsilon) \rightarrow 0$ .

**Exercise 9.** We want to generalise Exercise 6 of this sheet to the case when  $X_1, X_2, \dots$  take values in  $[0, +\infty[$ ; we still want to characterise the fact that  $\sum_{n=1}^{+\infty} X_n$  converges; again (ii) implies (iii).

1. Suppose (iii) again. Show that

$$\sum_{n=1}^{+\infty} \mathbb{P}(X_n > 1) < \infty.$$

(Hint: use Borel-Cantelli).

2. Suppose (iii). Show that

$$\sum_{n=1}^{\infty} \mathbb{E}[X_n \mathbf{1}_{X_n \leq 1}] < \infty.$$

(Use the previous exercise).

3. Suppose conversely that  $\sum_{n=1}^{+\infty} \mathbb{P}(X_n > 1) < \infty$  and  $\sum_{n=1}^{\infty} \mathbb{E}[X_n \mathbb{1}_{X_n \leq 1}] < \infty$ . Show that, almost surely, there is a finite number of  $n$ 's such that  $X_n > 1$ , and that  $\sum_{n \geq 1} X_n \mathbb{1}_{X_n \leq 1} < \infty$ . Conclude that (ii) is true.

**Exercise 10.** We want to improve on Exercise 7 in Sheet 6. Let  $\lambda > 0$  and let  $(X_n)_{n \geq 1}$  be a sequence of iid real random variables, with law  $\mathcal{E}(\lambda)$ . We already proved that the sequence

$$\frac{\max(X_1, \dots, X_n)}{\ln(n)}$$

converges to  $\frac{1}{\lambda}$  in probability. Let us prove that the convergence is, in fact, almost sure.

1. Prove that, for all  $\epsilon > 0$ , almost surely there exists only a finite number of  $n$ 's such that

$$\frac{\max(X_1, \dots, X_n)}{\ln(n)} \leq \frac{1}{\lambda} - \epsilon,$$

even though the sequence  $(\frac{\max(X_1, \dots, X_n)}{\ln(n)})_{n \geq 1}$  is not independent.

2. Show that the following statements are equivalent:

- Almost surely, the number of  $n$  such that

$$\frac{1}{\ln(n)} \max(X_1, \dots, X_n) \geq \frac{1}{\lambda} + \epsilon$$

is finite.

- Almost surely, the number of  $n$  such that

$$\frac{1}{\ln(n)} X_n \geq \frac{1}{\lambda} + \epsilon$$

is finite.

3. Deduce that the sequence  $(\frac{\max(X_1, \dots, X_n)}{\ln(n)})_{n \geq 1}$  converges almost surely to  $\frac{1}{\lambda}$ .

**Exercise 11.** Let  $(X_n)_{n \geq 1}$  be a sequence of iid, square-integrable real random variables, with expectation  $m$  and variance  $\sigma^2$ .

1. Show that for all  $\delta > 0$ , with probability at least  $1 - \delta$  one has

$$\left| \frac{1}{n} \sum_{i=1}^n X_i - m \right| \leq \frac{\sigma}{\sqrt{\delta n}}.$$

Deduce that  $\frac{1}{n} \sum_{i=1}^n X_i$  converges in probability to  $m$ .

2. Suppose now that  $X_n$  follow  $\mathbb{N}(m, \sigma)$ . Give a simple formula for the function  $M : \mathbb{R} \rightarrow \mathbb{R}$  defined by

$$M : u \mapsto \mathbb{E}[\exp(u(X_1 - m))].$$

3. Prove that, for every  $a \in \mathbb{R}$  and every  $u > 0$ ,

$$\mathbb{P}\left(\frac{1}{n} \sum_{i=1}^n X_i \geq m + \frac{a}{n}\right) \leq e^{-ua} (M(u))^n;$$

show a similar inequality for  $\mathbb{P}\left(\frac{1}{n} \sum_{i=1}^n X_i \leq m - \frac{a}{n}\right)$ .

4. Deduce that, for every  $\epsilon > 0$ ,

$$\mathbb{P}\left(\left|\frac{1}{n} \sum_{i=1}^n X_i - m\right| \geq \epsilon\right) \leq 2 \exp\left(\frac{-n\epsilon^2}{2\sigma^2}\right).$$

5. Find  $\delta$  such that  $\frac{\sigma}{\sqrt{\delta n}} = \epsilon$  and compare the above bound with the one in question 1.